

# Augmented Reality: A Computational Framework Applied to Modeling the Dynamics of Air Pollution

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## Abstract

In recent years, we have developed a new Augmented Reality (AR) framework to combine real data with computer-generated synthetic samples to “look under the hood”, as it were, for gaining insights into rare, dynamic phenomena. Using data fusion and density ratio model, AR allows us to estimate the tail probabilities of exceeding large thresholds that are far beyond the limited range of observations in moderately sized data. Such thresholds represent extreme events such as the drastic change in air pollution levels in Washington DC caused by lockdown due to the COVID-19 pandemic in 2020, as modeled in this study.

*Key words:* Data fusion; Tail probabilities; Density ratio model; Synthetic data; Air pollution.

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## 1. Introduction

In its February 4, 2017, edition, *The Economist* claimed that “Replacing the real world with a virtual one is a neat trick. Combining the two could be more useful.” Combining real data with synthetic data produces augmented reality (AR), which, we believe, opens up new perspectives regarding statistical inference. Indeed, augmentation of observations of the real world with virtual information is transforming engineering, healthcare and AI with emerging powerful technologies such as robotics, Internet of Things, and more recently, Digital Twins (Tao and Qi, 2019).

In a recent article, we advanced the notion of repeated AR in the estimation of very small tail probabilities even from moderately sized samples (Kedem and Pyne, 2021). Our approach, much like the bootstrap, is computationally intensive. However, unlike bootstrap, we look repeatedly outside the sample. Synthesis of a given sample of real world observations repeatedly with computer-generated data is based on repeated out of sample fusion (ROSF, Kedem *et al.* 2019; Zhang, Pyne and Kedem, 2020). This strategy proves to be useful for inference in various surveillance applications in which the available datasets usually have a limited range of observations and a moderate size due to limited storage capacity.

In particular, we are interested to estimate the tail probability  $p$  of observations exceeding a given high threshold  $T$ . Our repeated AR approach is based on numerous data fusions. We use an iterative method that can generate a large number of upper bounds  $B_i$  for

$p$ . Say, such a method is fast and probabilistic, and the upper bounds exceed  $p$  with a 95% chance. Thus, many of these exceed  $p$  but many do not. Therefore, there are subsequences of ordered upper bounds which approach  $p$  from above and from below. We showed how upper bounds can be produced by repeated fusion of real data with computer-generated samples, where the number of fusions is arbitrarily large, and where the support of the generated data is large enough so that it ranges beyond  $T$ . Hence, using the connection between the real and generated data, we have a computational approach to “peek” into the realm above  $T$ .

Notably, the repeated AR approach allows us to model many phenomena of sudden yet significant change that are of great interest to researchers, *e.g.*, for predicting stock market crashes, disease outbreaks, and extreme climatic events. On January 12, 2021, it was reported in the *New York Times* that “America’s greenhouse gas emissions from energy and industry plummeted more than 10 percent in 2020, reaching their lowest levels in at least three decades as the coronavirus pandemic slammed the brakes on the nation’s economy”. It pointed out that “transportation, the nation’s largest source of greenhouse gases, saw a 14.7 percent decline in emissions in 2020 as millions of people stopped driving to work” due to lockdowns that were implemented in many states of the U.S. over the course of the COVID-19 pandemic.

Nitrogen dioxide ( $\text{NO}_2$ ) is a gaseous pollutant emitted from the burning of fossil fuels at high temperatures primarily by vehicles, and thus, its level is a good indicator of traffic volume at a given area over a given interval of time. According to the American Lung Association,  $\text{NO}_2$  causes a range of harmful effects on the lungs, including increased inflammation of the airways, worsened cough and wheezing, reduced lung function, increased asthma attacks, and a greater likelihood of emergency department and hospital admissions. The U.S. Environmental Protection Agency’s (EPA) National Ambient Air Quality Standard (NAAQS), therefore, measures  $\text{NO}_2$  as an indicator for the  $\text{NO}_X$  family of air pollutants.

Given the sharp reduction in traffic after stay-at-home orders were enforced in many areas of the U.S., in this study, we are interested to model the resulting dynamics of air pollution at a given area. At the capital Washington DC, the stay-at-home order came into effect on April 1, 2020. To analyze the differences between the two periods, pre- and post-order, of 3 months on each side, we resort to two methods. First, we test for similarity in the levels of  $\text{NO}_2$  in the morning air in the two periods by using their respective probability distributions. This is done by fusion of data from the two periods as described in a previous study (Kedem *et al.*, 2017). Second, we estimate the tail probability of  $\text{NO}_2$  level exceeding  $T = 100$  parts per billion (ppb) in each of the two periods. This is done by repeated fusion of the data with computer generated samples (Kedem *et al.*, 2019, Kedem and Pyne, 2020).

## 2. Data and Methods

### 2.1. Air Pollution Data

The  $\text{NO}_2$  emissions data were collected at four monitoring stations of the U.S. Environmental Protection Agency (EPA) in Washington DC area, for the pre- and post-lockdown periods of January-March and April-June, 2020. In this study, we focused on the morning readings, *i.e.*, the hourly surface levels of  $\text{NO}_2$  between 6 am and 9 am. For each period, a random sample of size 200 was selected from the data collected at the locations with the (latitude, longitude) coordinates of (38.895572,  $-76.958072$ ), (38.921847,  $-77.013178$ ), (38.970092,  $-77.016715$ ), and (38.89477,  $-76.953426$ ). Thus, we obtained a  $\text{NO}_2$  sample of size 200 from the first period (January 1–March 31), and another sample of size 200 from the second period (April 1–June 30).

## 2.2. Testing for equidistribution

Let  $X_0$  be a sample of  $\text{NO}_2$  observations of size 200 from January-March, following an unknown probability density (pdf)  $g(x)$ ,  $x \in (0, \infty)$ , and let  $G(x)$  denote the corresponding unknown distribution function (CDF). Similarly, let  $X_1 \sim g_1, G_1$  be a sample of size 200 from the second period of April-June, following unknown pdf  $g_1(x)$  and CDF  $G_1(x)$ ,  $x \in (0, \infty)$ . We assume the *density ratio model* (Qin and Zhang 1997, Lu 2007)

$$\frac{g_1(x)}{g(x)} = \exp(\alpha_1 + \beta_1' h(x)) \quad (1)$$

where  $\alpha_1$  is a scalar parameter,  $\beta_1$  is a  $2 \times 1$  vector parameter, and  $h(x) = (x \log x)$ . We now combine or fuse the two samples and estimate the parameters in (1) from the combined sample of size of 400. Kernel density estimates of  $g, g_1$  are shown in Figure 1. From the fits in Figure 1 (bottom panel), we see that the estimated  $g, g_1$  are very close to the corresponding histograms, indicating that the choice of “gamma tilt”  $h(x) = (x \log x)$  is sensible.

## 2.3. Estimation of tail probabilities

The basic idea here is to fuse each  $\text{NO}_2$  sample with numerous computer-generated “synthetic” samples. This strategy is referred to as *repeated out of sample fusion* (ROSF in Kedem *et al.*, 2019) or *repeated augmented reality* (repeated AR in Kedem and Pyne, 2020).

If  $p = P(X > T)$  is a tail probability to be estimated, we generate numerous upper bounds  $B$ 's for  $p$  where most are above  $p$  but many are below  $p$ . If  $B_{(j)}$  are the corresponding order statistics, then there are  $B_{(j)}$  which bound  $p$  from above and there are  $B_{(j)}$  which bound  $p$  from below, yet some  $B_{(j)}$  fall in a small neighborhood of  $p$ . In this paper, we generated 10,000 such  $B$ 's. The problem is to find  $B_{(j)}$  in a small neighborhood of  $p$ . This is addressed by an iterative algorithm which produces subsequences of the  $B_{(j)}$  sequence which converge to a small neighborhood of  $p$  from above and from below as follows:

$$B_{(j^1)} < B_{(j^2)} < B_{(j^3)} < \cdots < B_{(j^n)} < p < B_{(j_m)} < \cdots < B_{(j_3)} < B_{(j_2)} < B_{(j_1)}$$

where  $B_{(j^n)}$  and  $B_{(j_m)}$  are very close to  $p$ .

We now have two relationships. For a sufficiently large number of fusions, say 10,000, there are  $B_{(j)}$  which approach  $p$  from above and from below, so that there is a  $B_{(j)}$  closest to  $p$ . This establishes a relationship between  $B_{(j)}$  and  $p$ .

With  $N = 1000$  (see the remark below), another relationship between  $B_{(j)}$  and  $p$  is obtained from the well known distribution of order statistics,

$$P(B_{(j)} > p) = \sum_{k=0}^{j-1} \binom{N}{k} [F_B(p)]^k [1 - F_B(p)]^{N-k} \quad (2)$$

where  $F_B$  is the distribution of  $B_i$  (not  $B_{(j)}$ ), which can be computed since  $F_B$  practically coincides with the empirical distribution of  $B_1, \dots, B_{10,000}$ .

The iterative algorithm consists of the following steps, starting from some  $j$ .

Step 1:

From (2) we can get the smallest  $p_j$  such that

$$P(B_{(j)} > p_j) = \sum_{k=0}^{j-1} \binom{N}{k} [F_B(p_j)]^k [1 - F_B(p_j)]^{N-k} \leq 0.95, \quad (3)$$

Step 2:

From Step 1 we get a  $j$  corresponding to the smallest  $p_j$ . Use this  $j$  and go back to Step 1.

Convergence is reached when for some  $k$ , we get the same probability values  $p_{j_k}$ . For further details of the algorithm, see Kedem and Pyne (2020).

Remark: We get 10,000 upper bounds  $B_1, \dots, B_{10,000}$  from which  $F_B$  is obtained. However, due to computational limitations, in (2) we use  $N = 1000$ .

$$m=2, h(x)=(x, \log x)$$

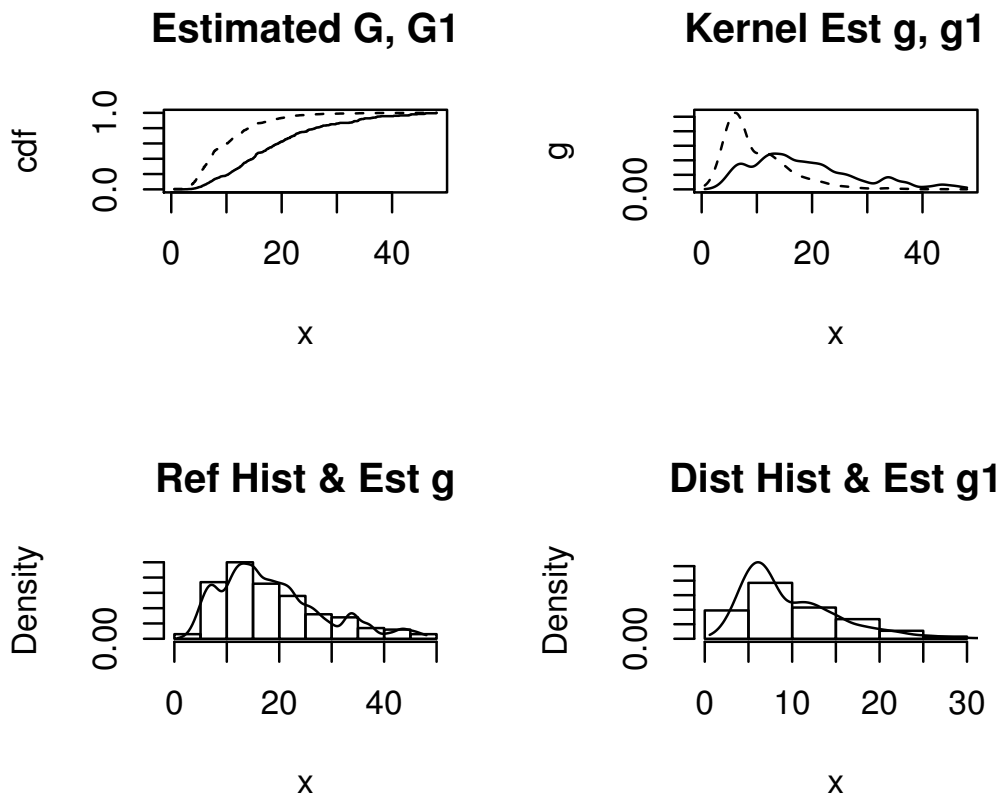


Figure 1: January-March (solid) vs. April-March (dashed) NO<sub>2</sub> distributions

### 3. Results

The likelihood ratio test of equidistribution  $H_0 : \beta_1 = 0$  gives a  $p$ -value of 0, indicating that the behavior in the two periods is completely different. This is also seen graphically from the plots of the two CDF's in in Figure 1 (top panel). We see that  $\hat{G}$  (solid line) is shifted much to the right relative to  $\hat{G}_1$  (dashed line), indicating a great reduction in NO<sub>2</sub> levels in the second period of April-June relative to the first period of January-March.

Indeed a 95% confidence interval for the NO<sub>2</sub> mean in January-March is approximately (14.14,17.24) whereas the same for April-June is approximately (9.11,10.69), again indicating a great reduction in NO<sub>2</sub> levels in latter period.

By setting the threshold  $T = 100$ , we estimated  $p = P(X > 100)$ , the probability that NO<sub>2</sub> exceeds a level of  $T = 100$ , for the two different (pre- and post-shutdown) periods in 2020. Recall that  $X_0$  is a sample of size 200 from the period of January-March. Fusing  $X_0$  10,000 times with generated Uniform(0,180) samples, the algorithm after one iteration gave:

$$B_{(810)} \rightarrow 0.00016 \rightarrow B_{(808)} \rightarrow 0.00016 = \hat{p} = 0.00016 \leftarrow B_{(809)} \leftarrow 0.00016 \leftarrow B_{(813)}$$

Thus, for the period of January-March 2020, we get  $\hat{p} = 0.00016$ .

Recall that  $X_1$  is a sample of size 200 from the period of April-June. Again, fusing  $X_1$  10,000 times with generated Uniform(0,180) samples, the algorithm gave after eight iterations going down, and a single iteration going up:

$$B_{(500)} \rightarrow 6.2e - 07 \rightarrow B_{(680)} \rightarrow 6.2e - 07 = \hat{p} = 6.2e - 07 \leftarrow B_{(690)} \cdots \leftarrow 2.5e - 05 \leftarrow B_{(900)}$$

Thus for the period April-June 2020, we get  $\hat{p} = 0.00000062$ , which is much smaller than  $\hat{p} = 0.00016$  from January-March 2020, echoing the previous results that the NO<sub>2</sub> levels had, in comparison, decreased significantly during April-June 2020 in Washington DC.

#### 4. Discussion

The COVID-19 pandemic has highlighted the need for systematic monitoring and rigorous modeling of dynamic phenomena that can exact a high toll in the form of human suffering and rapid losses in various sectors such as breakdown of supply chains and reduced mobility. Similar lessons are learnt from other areas including extreme climatic events and sudden crashes in the markets. In public health, surveillance is conducted routinely to guard against disease outbreaks and environmental exposures. In this study, we demonstrated how the repeated AR approach could provide a computational framework for modeling the dynamics of air pollution due to traffic emissions during a period of sudden, sharp change.

While estimation of small tail probabilities has long been a topic of research, many of the commonly used methods rely on large number of observations, which makes them less practical for modeling of dynamic phenomena. Methods such as peaks-over-threshold (POT) require observations beyond a threshold, whereas block maxima (BM) require sufficient data such that maxima from each block can be used for estimation. In comparison, both the availability as well as the reliability of computer-generated samples that are representative of real data are increasing with the development of new, powerful computational platforms, e.g., generative adversarial networks (GANs), thus allowing for easier data augmentation.

In this study, we built our AR approach on a density ratio model that starts with a common reference distribution for all sources of information, and then models the individual distributions as *distortions* (e.g., gamma tilt) of that “baseline”. Further, we estimate very small tail probabilities even from moderately sized samples. Fusing these repeatedly with computer-generated synthetic data is particularly insightful when the data at hand falls short of the high threshold of interest. Our iterative algorithm constructs a “*B*-sequence” of bounds that contains a point whose ordinate is very close to the target tail probability, as ensured by the Glivenko-Cantelli theorem. For further details as well as the strengths and the limitations of the repeated AR approach, the reader is referred to Kedem and Pyne (2020) and Kedem *et al.* (2021).

A limitation of the present study is that it does not explicitly account for the fact that the level of NO<sub>2</sub> typically shows a decrease with the advent of spring and summer as it

dissociates in sunlight, and tends to collect less near the surface during that period. While the shift in  $\text{NO}_2$  was much larger in 2020 compared to previous years, a multi-year extension of our model would be more insightful. Since the primary aim of this study is to introduce our AR computational framework, we plan to address this in our future work.

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